



L_1 -consistent estimation of the density of residuals in random design regression models

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ABSTRACT

In this paper we study the problem of estimating the density of the error distribution in a random design regression model, where the error is assumed to be independent of the design variable. Our main result is that the L_1 error of the kernel density estimate applied to residuals of a consistent regression estimate converges with probability 1 to 0 regardless of the form of the true density. We demonstrate that this result is in general no longer true if the error distribution and the design variable are dependent.

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1. Introduction

In this paper we study the problem of estimating the density of the error distribution in a random design regression model. More precisely, we assume that we have given data

$$\mathcal{D}_n = \{(X_1, Y_1), \dots, (X_n, Y_n)\},$$

where

$$Y_i = m(X_i) + \epsilon_i$$

for some $m: \mathbb{R}^d \rightarrow \mathbb{R}$, $\mathbb{R}^d$ -valued random (design) variables $X_1, \dots, X_n$ and real-valued random variables $\epsilon_1, \dots, \epsilon_n$ with zero expectation. We assume that $(X_1, \epsilon_1), (X_2, \epsilon_2), \dots$ are independent and identically distributed and that X_1, ϵ_1 are independent. We let X be a random variable distributed as X_1 and ϵ be a random variable with density f with respect to Lebesgue measure. It is assumed that all ϵ_i 's are distributed as ϵ . We use $\mathbf{P}_X$ for the probability law of a random variable X . We consider the problem of estimating f from the data $\mathcal{D}_n$.

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Estimating the density of the error distribution in nonparametric regression models has been dealt with by several researchers. Ahmad (1992) showed that under a Lipschitz condition for the kernel function, the Parzen–Rosenblatt density estimator (Parzen, 1962; Rosenblatt, 1956) converges in probability at every continuity point to the real density of the residuals. In the case of a continuous error density, the same estimator is pointwise and uniformly consistent (Cheng, 2004), and, in addition, the histogram error density estimator is uniformly and in L_1 consistent (Cheng, 2002). Efromovich (2005) investigated using a homoscedastic regression model estimates which are as good as estimates using an oracle that knows the underlying regression errors. Using the heteroscedastic nonparametric regression model, where the Y_i 's have different variances, Efromovich (2006) generalized his optimal estimation for a twice-differentiable error density with finite support. Estimators of the residual distribution function include that of Akritas and Van Keilegom (2001), who extended the results of Durbin (1973) and Loynes (1980) to a weak convergence result for a distribution function estimator in a nonparametric heteroscedastic regression model. The empirical distribution function of residuals was used as an estimator in a heteroscedastic model with multivariate covariates by Neumeyer and Van Keilegom (2010). For general results in density estimation we refer the reader to the books of Devroye and Györfi (1985), Devroye (1987) and Devroye and Lugosi (2000), and to the article (Mnatsakanov and Khmaladze, 1981), where the necessary and sufficient conditions for L_1 -consistency of general density estimates are derived.

In this paper the main aim is to derive L_1 -consistent estimates of f . This is important, because Scheffé's Lemma implies that the L_1 error of a density estimate equals twice the total variation distance (see, e.g., Devroye and Györfi (1985)) and hence an L_1 -consistent density estimate allows simultaneous estimation of all probabilities. In order to estimate the density of the error distribution, we split the data in two parts, use the first part to compute a regression estimate, compute its residuals on the second part and use them as data for a standard kernel density estimate. We show that the resulting density estimate is strongly consistent in L_1 for all densities f provided that we use in the first step a consistent regression estimate. Furthermore, we show that this result no longer holds in the case where the design variable and the errors in the regression model are dependent. The estimates are defined in Section 2, the main results are presented in Section 3, and Section 4 contains the proofs.

2. Definition of the estimates

We start by splitting the sample in two parts: a first part of size n' and the second part containing the $n'' = n - n'$ remaining data points. We assume that we have given a regression estimate

$$m_{n'}(\cdot) = m_{n'}(\cdot, \mathcal{D}_{n'}) : \mathbb{R}^d \rightarrow \mathbb{R}$$

satisfying

$$\int |m_{n'}(x) - m(x)| \mathbf{P}_X(dx) \rightarrow 0 \quad \text{a.s.}$$

(for $n \rightarrow \infty$), which we apply to the first part of the data. We use $m_{n'}$ to compute the residuals on the second part of the data, i.e., we set

$$\hat{\epsilon}_j = Y_{n'+j} - m_{n'}(X_{n'+j}) \quad (j = 1, \dots, n'')$$

and we use them to compute the Parzen–Rosenblatt kernel density estimate

$$f_n(x) = \frac{1}{n''h_{n''}} \sum_{j=1}^{n''} K\left(\frac{x - \hat{\epsilon}_j}{h_{n''}}\right).$$

Here $h_{n''} > 0$ is the bandwidth, and $K: \mathbb{R} \rightarrow \mathbb{R}_+$ is the kernel function, which we assume to be a density (in general, only integrability to 1 is needed, but we will be happy with the subclass of densities).

3. The main results

Our first result deals with the consistency of f_n .

Theorem 1. Let $X_1, X_2, \dots$ be i.i.d. $\mathbb{R}^d$ -valued random variables distributed as X , and let $\epsilon_1, \epsilon_2, \dots$ be i.i.d. real-valued random variables with $\mathbf{E}\{\epsilon_1\} = 0$, and having common density f , and assume that both sequences are also independent of each other. Define

$$Y_i = m(X_i) + \epsilon_i \quad (i \in \mathbb{N})$$

for some function $m: \mathbb{R}^d \rightarrow \mathbb{R}$, and let the estimate f_n be defined as in Section 2. Assume that K is a density with compact support satisfying

$$\int_{\mathbb{R}} K^2(u) du < \infty, \tag{1}$$

such that

$$h_n \rightarrow 0 \quad (n \rightarrow \infty) \quad \text{and} \quad n \cdot h_n \rightarrow \infty \quad (n \rightarrow \infty) \quad (2)$$

and that the regression estimate m_n satisfies

$$\int_{\mathbb{R}^d} |m_n(x) - m(x)| \mathbf{P}_{X_1}(dx) \rightarrow 0 \quad \text{a.s.} \quad (3)$$

Finally, assume that both $n' \rightarrow \infty$ and $n'' \rightarrow \infty$ as $n \rightarrow \infty$. Then

$$\int_{\mathbb{R}} |f_n(x) - f(x)| dx \rightarrow 0 \quad \text{a.s.}$$

Remark 1. The L_1 error of the above density estimate tends to zero regardless of the form of the density of ϵ_1 , provided the regression estimate is $L_1(\mathbf{P}_{X_1})$ -consistent. By Cauchy–Schwarz, (3) is implied by

$$\int_{\mathbb{R}^d} |m_n(x) - m(x)|^2 \mathbf{P}_{X_1}(dx) \rightarrow 0 \quad \text{a.s.}, \quad (4)$$

and if we assume $\mathbf{E}\{Y_1^2\} < \infty$, there are many different estimates which are universally consistent in the sense that (4) holds for all distributions; cf., e.g., Devroye et al. (1994), Györfi and Walk (1996, 1997), Kohler and Krzyżak (2001), Lugosi and Zeger (1995), Nobel (1996), Walk (2002), or Györfi et al. (2002) and the literature cited therein. So under this additional assumption our estimate of the density of the error distribution in our regression model is strongly consistent in L_1 for all densities.

In Theorem 1, the generic pair (X, ϵ) is independent. Without this independence condition, one has to be much more careful. Noting that Theorem 1 was formulated in terms of a very general L_1 -consistent regression estimate (3), the lack of independence will force one to at least make a specific choice of regression function estimate. This is captured in Theorem 2.

Theorem 2. Let $d = 1$. Assume that $|n' - n''| \leq 1$, and $n = n' + n''$. There exists a regression function m and a distribution of (X, ϵ) with $\mathbf{E}\{\epsilon|X\} = 0$, such that ϵ has a density f (with respect to the Lebesgue measure), and such that for any sequence of bandwidths h_n satisfying (2) there exists a regression estimate m_n satisfying (3) with the property that the corresponding density estimate f_n from Section 2 in the case of the naive kernel (i.e., the kernel $K(x) = (1/2)I_{[-1,1]}(x)$) is such that

$$\limsup_{n \rightarrow \infty} \int_{\mathbb{R}} |f_n(x) - f(x)| dx \geq 1.$$

Remark 2. Theorem 1 showed that when the pair (X, ϵ) is independent, the density estimate of Section 2 is strongly L_1 -consistent whenever the regression estimate is strongly $L_1(\mathbf{P}_{X_1})$ -consistent. By Theorem 2 this is no longer true in all generality if we omit the independence assumption. The counterexample in Theorem 2 has two special properties—first of all, conditional on X , ϵ is a shifted Bernoulli random variable without a density, but its marginal distribution has a density. Secondly, and more importantly, the example uses a special regression estimate that one would not encounter in statistical practice. It is still an open problem whether there exists any regression estimate such that the corresponding density estimate from Section 2 is strongly L_1 -consistent for all distributions of (X, Y) where $\epsilon = Y - \mathbf{E}\{Y|X\}$ has a density f and where $\mathbf{E}|Y| < \infty$.

4. Proofs

4.1. Proof of Theorem 1

First we show that

$$\int_{\mathbb{R}} |f_n(x) - f(x)| dx - \mathbf{E} \left\{ \int_{\mathbb{R}} |f_n(x) - f(x)| dx \middle| \mathcal{D}_{n'} \right\} \rightarrow 0 \quad (5)$$

almost surely (applied conditioned on $\mathcal{D}_{n'}$).

Define

$$g(\hat{\epsilon}_1, \dots, \hat{\epsilon}_{n''}) := \int_{\mathbb{R}} \left| \frac{1}{n'' \cdot h_{n''}} \sum_{j=1}^{n''} K\left(\frac{x - \hat{\epsilon}_j}{h_{n''}}\right) - f(x) \right| dx.$$

Scheffé's Lemma implies

$$\begin{aligned} |g(\hat{\epsilon}_1, \dots, \hat{\epsilon}_{n''}) - g(\hat{\epsilon}_1, \dots, \hat{\epsilon}_{i-1}, \hat{\epsilon}'_i, \hat{\epsilon}_{i+1}, \hat{\epsilon}_{n''})| &\leq \int_{\mathbb{R}} \left| \frac{1}{h_{n''}} K\left(\frac{x - \hat{\epsilon}_i}{h_{n''}}\right) - \frac{1}{h_{n''}} K\left(\frac{x - \hat{\epsilon}'_i}{h_{n''}}\right) \right| dx \\ &= 2 \cdot \int_{\mathbb{R}} \left(\frac{1}{h_{n''}} K\left(\frac{x - \hat{\epsilon}_i}{h_{n''}}\right) - \frac{1}{h_{n''}} K\left(\frac{x - \hat{\epsilon}'_i}{h_{n''}}\right) \right)_+ dx \\ &\leq 2. \end{aligned}$$

From McDiarmid's inequality (McDiarmid (1989); see also Devroye (1991) and Theorem A.2. in Györfi et al. (2002)), we see that

$$\sum_{n=1}^{\infty} \mathbf{P} \left[\left| \int_{\mathbb{R}} |f_n(x) - f(x)| dx - \mathbf{E} \left\{ \int_{\mathbb{R}} |f_n(x) - f(x)| dx \middle| \mathcal{D}_{n'} \right\} \right| \geq \epsilon \middle| \mathcal{D}_{n'} \right] \leq \sum_{n=1}^{\infty} \exp \left(\frac{-n'' \epsilon^2}{2} \right) < \infty.$$

We obtain (5) by an application of the lemma of Borel and Cantelli.

Next we show that

$$\mathbf{E} \left\{ \int_{\mathbb{R}} |f_n(x) - f(x)| dx \middle| \mathcal{D}_{n'} \right\} \rightarrow 0 \quad \text{a.s.}$$

Scheffé's Lemma implies that

$$\int_{\mathbb{R}} |f_n(x) - f(x)| dx = 2 \cdot \int_{\mathbb{R}} (f(x) - f_n(x))_+ dx \leq 2 \cdot \int_B |f(x) - f_n(x)| dx + 2 \cdot \int_{B^c} f(x) dx.$$

Hence it suffices to show

$$\mathbf{E} \left\{ \int_B |f_n(x) - f(x)| dx \middle| \mathcal{D}_{n'} \right\} \rightarrow 0 \quad \text{a.s.}$$

for any compact set $B \subseteq \mathbb{R}$. Let B be an arbitrary compact set in $\mathbb{R}$. Set

$$f_n^*(x) = \frac{1}{n'' \cdot h_{n''}} \sum_{j=1}^{n''} \mathbf{E} \left\{ K\left(\frac{x - \hat{\epsilon}_j}{h_{n''}}\right) \middle| \mathcal{D}_{n'} \right\}.$$

Then

$$\mathbf{E} \left\{ \int_B |f_n(x) - f(x)| dx \middle| \mathcal{D}_{n'} \right\} \leq \mathbf{E} \left\{ \int_B |f_n(x) - f_n^*(x)| dx \middle| \mathcal{D}_{n'} \right\} + \mathbf{E} \left\{ \int_B |f_n^*(x) - f(x)| dx \middle| \mathcal{D}_{n'} \right\}.$$

In the first step of the proof we show that

$$\mathbf{E} \left\{ \int_B |f_n(x) - f_n^*(x)| dx \middle| \mathcal{D}_{n'} \right\} \rightarrow 0 \quad (n \rightarrow \infty).$$

By Cauchy-Schwarz and the inequality of Jensen we have

$$\begin{aligned} \mathbf{E} \left\{ \int_B |f_n(x) - f_n^*(x)| dx \middle| \mathcal{D}_{n'} \right\} &\leq \mathbf{E} \left\{ \sqrt{\int_B 1 dx} \cdot \sqrt{\int_B |f_n(x) - f_n^*(x)|^2 dx} \middle| \mathcal{D}_{n'} \right\} \\ &\leq \sqrt{\int_B 1 dx} \cdot \sqrt{\mathbf{E} \left\{ \int_B |f_n(x) - f_n^*(x)|^2 dx \middle| \mathcal{D}_{n'} \right\}}. \end{aligned}$$

Now

$$\begin{aligned} \mathbf{E} \left\{ \int_B |f_n(x) - f_n^*(x)|^2 dx \middle| \mathcal{D}_{n'} \right\} &= \int_B \mathbf{E} \left\{ \left| \frac{1}{n'' \cdot h_{n''}} \sum_{j=1}^{n''} \left(K\left(\frac{x - \hat{\epsilon}_j}{h_{n''}}\right) - \mathbf{E} \left\{ K\left(\frac{x - \hat{\epsilon}_j}{h_{n''}}\right) \middle| \mathcal{D}_{n'} \right\} \right) \right|^2 \middle| \mathcal{D}_{n'} \right\} dx \\ &\leq \frac{1}{n''^2 \cdot h_{n''}^2} \sum_{j=1}^{n''} \int_B \mathbf{E} \left\{ K^2\left(\frac{x - \hat{\epsilon}_j}{h_{n''}}\right) \middle| \mathcal{D}_{n'} \right\} dx \\ &= \frac{1}{n''^2 \cdot h_{n''}^2} \sum_{j=1}^{n''} \int_B \int_{\mathbb{R}} K^2\left(\frac{x - u - m(X_{n'+j}) + m_{n'}(X_{n'+j})}{h_{n''}}\right) f(u) du dx \\ &= \frac{1}{n''^2 \cdot h_{n''}^2} \sum_{j=1}^{n''} \int_{\mathbb{R}} f(u) \int_B \frac{1}{h_{n''}} K^2\left(\frac{x - u - m(X_{n'+j}) + m_{n'}(X_{n'+j})}{h_{n''}}\right) dx du \\ &\leq \frac{1}{n'' \cdot h_{n''}} \int_{\mathbb{R}} K^2(z) dz \rightarrow 0 \quad (n \rightarrow \infty). \end{aligned}$$

In the second step of the proof we show that

$$\mathbf{E} \int_{\mathbb{R}} \left| f_n^*(x) - \mathbf{E} \left\{ \frac{1}{h_{n''}} K \left(\frac{x - \epsilon_1}{h_{n''}} \right) \right\} \right| dx \rightarrow 0 \quad (n \rightarrow \infty).$$

To do this we observe that

$$\begin{aligned} \int_{\mathbb{R}} \left| f_n^*(x) - \mathbf{E} \left\{ \frac{1}{h_{n''}} K \left(\frac{x - \epsilon_1}{h_{n''}} \right) \right\} \right| dx &= \int_{\mathbb{R}} \left| \frac{1}{n''} \sum_{j=1}^{n''} \int_{\mathbb{R}} \frac{1}{h_{n''}} K \left(\frac{x - u + m_{n'}(X_{n'+j}) - m(X_{n'+j})}{h_{n''}} \right) f(u) du \right. \\ &\quad \left. - \int_{\mathbb{R}} \frac{1}{h_{n''}} K \left(\frac{x - u}{h_{n''}} \right) f(u) du \right| dx \\ &\leq \int_{\mathbb{R}} \frac{1}{n''} \sum_{j=1}^{n''} \int_{\mathbb{R}} \frac{1}{h_{n''}} K \left(\frac{x - z}{h_{n''}} \right) |f(z - m(X_{n'+j}) + m_{n'}(X_{n'+j})) - f(z)| dz dx \\ &\leq \frac{1}{n''} \sum_{j=1}^{n''} \int_{\mathbb{R}} \int_{\mathbb{R}} \frac{1}{h_{n''}} K \left(\frac{x - z}{h_{n''}} \right) \cdot |f(z + m_{n'}(X_{n'+j}) - m(X_{n'+j})) \\ &\quad - g_f(z + m_{n'}(X_{n'+j}) - m(X_{n'+j}))| dz dx \\ &\quad + \frac{1}{n''} \sum_{j=1}^{n''} \int_{\mathbb{R}} \int_{\mathbb{R}} \frac{1}{h_{n''}} K \left(\frac{x - z}{h_{n''}} \right) \\ &\quad \times |g_f(z + m_{n'}(X_{n'+j}) - m(X_{n'+j})) - g_f(z)| dz dx \\ &\quad + \frac{1}{n''} \sum_{j=1}^{n''} \int_{\mathbb{R}} \int_{\mathbb{R}} \frac{1}{h_{n''}} K \left(\frac{x - z}{h_{n''}} \right) \cdot |g_f(z) - f(z)| dz dx \\ &=: T_{1,n} + T_{2,n} + T_{3,n} \end{aligned}$$

with an arbitrary density $g_f: \mathbb{R}^d \rightarrow \mathbb{R}_+$.

Application of Fubini's theorem yields

$$\begin{aligned} T_{1,n} &= \frac{1}{n''} \sum_{j=1}^{n''} \int_{\mathbb{R}} |f(z + m_{n'}(X_{n'+j}) - m(X_{n'+j})) - g_f(z + m_{n'}(X_{n'+j}) - m(X_{n'+j}))| \cdot \int_{\mathbb{R}} \frac{1}{h_{n''}} K \left(\frac{x - z}{h_{n''}} \right) dx dz \\ &= \int_{\mathbb{R}} |f(z) - g_f(z)| dz \end{aligned}$$

and

$$\begin{aligned} T_{3,n} &= \frac{1}{n''} \sum_{j=1}^{n''} \int_{\mathbb{R}} |g_f(z) - f(z)| \cdot \int_{\mathbb{R}} \frac{1}{h_{n''}} K \left(\frac{x - z}{h_{n''}} \right) dx dz \\ &= \int_{\mathbb{R}} |f(z) - g_f(z)| dz. \end{aligned}$$

Choosing g_f as a density which approximates f in L_1 , both terms $T_{1,n}$ and $T_{3,n}$ can be made arbitrary small.

Assume that g_f is Lipschitz-continuous with Lipschitz constant $L > 0$ and let C be an arbitrary compact set in $\mathbb{R}$. By Scheffé's Lemma,

$$\begin{aligned} T_{2,n} &= \frac{1}{n''} \sum_{j=1}^{n''} \int_{\mathbb{R}} |g_f(z + m_{n'}(X_{n'+j}) - m(X_{n'+j})) - g_f(z)| \cdot \int_{\mathbb{R}} \frac{1}{h_{n''}} K \left(\frac{x - z}{h_{n''}} \right) dx dz \\ &= \frac{1}{n''} \sum_{j=1}^{n''} 2 \cdot \int_{\mathbb{R}} (g_f(z) - g_f(z + m_{n'}(X_{n'+j}) - m(X_{n'+j})))_+ dz \\ &\leq 2 \cdot \frac{1}{n''} \sum_{j=1}^{n''} \int_C |g_f(z) - g_f(z + m_{n'}(X_{n'+j}) - m(X_{n'+j}))| dz + 2 \cdot \int_{C^c} g_f(z) dz \\ &\leq 2 \cdot L \cdot \frac{1}{n''} \sum_{j=1}^{n''} |m_{n'}(X_{n'+j}) - m(X_{n'+j})| \int_C 1 dz + 2 \cdot \int_{C^c} g_f(z) dz. \end{aligned}$$

So

$$\limsup_{n \rightarrow \infty} \mathbf{E} \{T_{2,n} | \mathcal{D}_{n'}\} \leq 2L \cdot \int_{\mathbb{C}} 1 \, dz \cdot \lim_{n \rightarrow \infty} \int_{\mathbb{R}^d} |m_{n'}(x) - m(x)| \mathbf{P}_{X_1}(dx) + 2 \int_{\mathbb{C}^c} g_f(z) \, dz$$

and with $C \uparrow \mathbb{R}$ we get

$$\mathbf{E} \{T_{2,n} | \mathcal{D}_{n'}\} \rightarrow 0 \quad \text{a.s.}$$

From Devroye and Györfi (1985), Chapter 3, Theorem 1 we know that

$$\int_{\mathbb{R}} \left| \frac{1}{h_{n''}} \mathbf{E} \left\{ K \left(\frac{x - \epsilon_1}{h_{n''}} \right) \right\} - f(x) \right| dx \rightarrow 0 \quad (n \rightarrow \infty)$$

and the assertion is proved. $\square$

4.2. Proof of Theorem 2

Without loss of generality, we assume that n is even, so $n' = n'' = n/2$. Let $i \in \{1, \dots, n\}$ and $j \in \{1, \dots, n''\}$. Let X_i be i.i.d. and uniformly distributed on $[0, 1]$, and define

$$Y_i = \begin{cases} X_i & \text{with probability } \frac{1}{2} \\ -X_i & \text{with probability } \frac{1}{2}. \end{cases}$$

Then Y_i is uniformly distributed on $[-1, 1]$, $\mathbf{E}\{Y_i | X_i\} = 0$ and $\epsilon_i = Y_i - \mathbf{E}\{Y_i | X_i\} = Y_i$. Define

$$m_{n'}(x) = x - x'$$

with $x' = \lfloor \frac{x}{\sqrt{h_{n'}}} \rfloor \cdot \sqrt{h_{n'}} = \lfloor \frac{x}{\sqrt{h_{n''}}} \rfloor \cdot \sqrt{h_{n''}}$. Since

$$|m_{n'}(x)| \leq \sqrt{h_{n'}} \rightarrow 0 \quad (n \rightarrow \infty)$$

assumption (3) is valid. Estimating the residuals as described in Section 2 we get

$$\hat{\epsilon}_j = Y_{n'+j} - m_{n'}(X_{n'+j}) = \begin{cases} X'_{n'+j} & \text{with probability } \frac{1}{2} \\ -2X_{n'+j} + X'_{n'+j} & \text{with probability } \frac{1}{2}, \end{cases}$$

where $X'_{n'+j} = \lfloor \frac{X_{n'+j}}{\sqrt{h_{n''}}} \rfloor \cdot \sqrt{h_{n''}}$. Observe that $-2X_{n'+j} + X'_{n'+j} \leq 0$. Define

$$A_n := \bigcup_{k=0}^{\infty} [k\sqrt{h_n} - h_n, k\sqrt{h_n} + h_n].$$

Note that these intervals are disjoint if $h_n < \frac{1}{4}$. So there exists a natural number $N \in \mathbb{N}$ such that for every $n'' > N$, we have

$$\begin{aligned} \int_{[0,1] \cap A_{n''}^c} f(x) dx &= \sum_{k=0}^{\ell_{n''}-1} \int_{k\sqrt{h_{n''}}+h_{n''}}^{(k+1)\sqrt{h_{n''}}-h_{n''}} \frac{1}{2} dx \\ &= \frac{1}{2} \ell_{n''} (\sqrt{h_{n''}} - 2h_{n''}) \end{aligned}$$

with $\ell_{n''} := \max\{k \in \mathbb{N} : k\sqrt{h_{n''}} \leq 1\} = \lfloor \frac{1}{\sqrt{h_{n''}}} \rfloor$.

If $x \in A_{n''}^c \cap [0, 1]^d$ then $|x - k\sqrt{h_{n''}}| > h_{n''}$ for every $k \in \mathbb{N}$ and consequently $|x - \hat{\epsilon}_j| > h_{n''}$ for all $j \in \mathbb{N}$. Using the naive kernel in the definition of f_n , we have for every $n'' > N$,

$$\int_{[0,1] \cap A_{n''}^c} f_n(x) dx = \int_{[0,1] \cap A_{n''}^c} \frac{1}{2n''h_{n''}} \sum_{j=1}^{n''} \mathbf{1}_{\{|x - \hat{\epsilon}_j| \leq h_{n''}\}}(x) dx = 0.$$

By Scheffé's Lemma, we see that for every $n'' > N$,

$$\begin{aligned} \int |f_n(x) - f(x)| dx &= 2 \cdot \sup_{B \in \mathcal{B}} \left| \int_B f(x) dx - \int_B f_n(x) dx \right| \\ &\geq 2 \cdot \int_{[0,1] \cap A_{n''}^c} f(x) dx - 2 \cdot \int_{[0,1] \cap A_{n''}^c} f_n(x) dx \\ &= \ell_{n''} \sqrt{h_{n''}} - \ell_{n''} \sqrt{h_{n''}} \sqrt{h_{n''}} \\ &\rightarrow 1 \quad (n \rightarrow \infty) \end{aligned}$$

and the proof is complete. $\square$

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