ON THE POINTWISE AND THE INTEGRAL CONVERGENCE OF RECURSIVE KERNEL ESTIMATES OF PROBABILITY DENSITIES

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ABSTRACT. Let X_1, \ldots, X_n be independent R^d -valued random vectors with a common density f, and let f be estimated by $f_n(x) = n^{-1} \sum_i h_i^{-d} K((x-X_i)/h_i)$, where $\{h_n\}$ is a sequence of positive numbers and K is a bounded density on R^d .

Several results related to the weak or the strong pointwise consistency of f are discussed and derived in the first part of the paper. In the second part, weak conditions on $\{h_n\}$ are given insuring that $\int |f_n(x)-f(x)| dx \stackrel{n}{\to} 0$ in probability (or with probability one) for all densities f.

1. Introduction.

Let X_1, X_2, \ldots be a sequence of independent identically distributed random vectors taking values in R^d with a common unknown density f. To estimate recursively the density for each $x \in R^d$ we consider the Wolverton-Wagner-Yamato estimate

(1)
$$f_n(x) = n^{-1} \sum_{i=1}^n h_i^{-d} K((x-X_i)/h_i)$$

(see [25-27]), where K is a given bounded probability density and $\{h_n\}$ is a sequence of positive numbers satisfying

$$h_{n} \stackrel{n}{\to} 0$$

and

(3)
$$n h_n^{\stackrel{d}{\rightarrow} \infty}.$$

The estimate f_n is a recursive version of the celebrated Parzen-Rosenblatt kernel estimate [20], [23]

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(4)
$$g_n(x) = n^{-1} \sum_{i=1}^{n} h_n^{-d} K((x-X_i)/h_n).$$

For surveys on nonparametric density estimation, the reader is referred to Wegman [24] and Revesz [22].

It is well known that $\,{\rm g}_n^{}\,$ is weakly pointwise consistent if (2) and (3) hold, if $\,{\rm f}\,$ is continuous at $\,{\rm x}\,$, and if

(5)
$$||\mathbf{x}||^{\mathbf{d}} K(\mathbf{x}) \to 0 \text{ as } ||\mathbf{x}|| \to \infty$$

(see [6] and [20]). If, to these conditions, we add

(6)
$$\sum_{n=1}^{\infty} e^{-\alpha n \hat{h}_{n}^{d}} < \infty \text{ for all } \alpha > 0 ,$$

then g_n is strongly pointwise consistent as well, that is, $g_n(x) \stackrel{\eta}{\to} f(x)$ with probability one (w.p.1) [14]. Limiting himself to bounded continuous f, to bounded Riemann-integrable K and to sequences $\{h_n\}$ satisfying $h_n \stackrel{\eta}{\to} 0$ and

(7)
$$(n+k) h_{n+k}^{d} \ge c n h_{n}^{d} \text{ (some } c > 0, all } n,k > 0),$$

Deheuvels [12] shows that $g_n(x) \stackrel{n}{\rightarrow} f(x)$ w.p.1 for all x if and only if

(8)
$$n h_n^d / log log n \stackrel{n}{\rightarrow} \infty .$$

Condition (8) is implied by

(9)
$$n h_n^d / log n \stackrel{n}{\rightarrow} \infty ,$$

which in turn implies (6) but is equivalent to (6) for sequences satisfying (7). However, in the absence of (7), neither (6) nor (9) imply Deheuvels' condition (8); e.g., let $\operatorname{nh}_n^d = \sqrt{n}$ except when $n=2^{k}$ for some integer k, in which case we take $\operatorname{nh}_n^d = \log\log n$.

In the first part of this paper we discuss and prove similar pointwise results for the Wolverton-Wagner-Yamato estimate (1). For other recent work on recursive density estimation, the reader is referred to the papers of Ahmad and Lin [2], Banon [3], Carroll [7],

Davies [8], Davies and Wegman [9], and Rejto and Revesz [21]. The trouble with a recursive estimate like (1) is that the contribution of $\mathbf{X_i}$ is weighted by $\mathbf{h_i^{-d}}$ and that the estimate is therefore sensitive to sudden decreases in the value of $\mathbf{h_n}$. Recently, Deheuvels [10-12] proposed a particularly elegant estimate

$$\bar{f}_n(x) = \left(\sum_{i=1}^n h_i^d\right)^{-1} \sum_{i=1}^n K((x-X_i)/h_i)$$

which shares with (1) and (4) that it is a density on \mathbb{R}^d and which has the nice property that all X_i are seemingly equally weighted. (At least, due to the boundedness of K, the difference $\overline{f}_n - \overline{f}_{n-1}$ will remain small even if h_n is much smaller than h_{n-1}, \ldots, h_1 .) For d=1, Deheuvels [10-11] shows that \overline{f}_n is strongly pointwise consistent if f and K are bounded densities, if f is continuous at f and if

(10)
$$h_n \stackrel{n}{\to} 0 \quad \text{and} \quad \sum_{n=1}^{\infty} h_n^d = \infty .$$

Pointwise consistency is usually of limited value to the statistician. Fortunately, it is true that for a sequence of densities f_n that is almost everywhere weakly (strongly) pointwise consistent, the in probability (w.p.1) convergence to 0 of $\int |f_n(x) - f(x)| \, \mathrm{d}x$ follows (Glick [16]); that is, f_n is weakly (strongly) consistent in L_1 . In the second part of this report a variety of weak conditions on $\{h_n\}$ are derived insuring the weak (strong) consistency in L_1 of f_n , g_n , and \overline{f}_n for all densities f (i.e., f is not required to be bounded or even to be almost everywhere continuous).

2. Main Results.

Let us first recall a pointwise consistency theorem for the Parzen-Rosenblatt estimate (see [6], [14], [20]).

THEOREM 1. Let $\{h_n\}$ satisfy (2), (3), let K be a bounded probability density, and let Condition A hold.

Condition A. One of the following is true:

- (i) x is a continuity point of f and (5) holds,
- (ii) x is a Lebesgue point of f and f is bounded on R^d ,
- (iii) x is a Lebesgue point of f and K has compact support.

Under these conditions, $g_n(x) \stackrel{n}{\to} f(x)$ in probability. If, in addition, (6) holds, then $g_n(x) \stackrel{n}{\to} f(x)$ w.p.1.

It is worth recalling that $\ x\$ is a Lebesgue point for $\ f\$ if

$$\rho^{-d} \int_{\left| \left| y-x \right| \right| \le \rho} \left| f(y) - f(x) \right| dy \to 0 \quad \text{as} \quad \rho \to 0 \ ,$$

that all continuity points are Lebesgue points, and that almost all points x are Lebesgue points. Using the techniques of proof of Theorem 1, it is shown in the Appendix that f_n converges pointwise to f under essentially the same conditions.

THEOREM 2. $f_n(x) \stackrel{n}{\to} f(x)$ in probability if $\{h_n\}$ satisfies (2), (3), if K is a bounded probability density, and if Condition A holds. If, additionally, (8) is true, then $f_n(x) \stackrel{n}{\to} f(x)$ w.p.1.

Proof of Theorem 2. The proof parallels in several respects the proof of Theorem 1 which can be found in [14]. The main tool we need is an inequality due to Bennett [4] for the sum of independent random variables Y_1, \ldots, Y_n with $E\{Y_i\} = 0$, $E\{Y_i^2\} \leq \sigma^2$, and $|Y_i| \leq b$:

$$P\left\{\left|n^{-1}\sum_{i=1}^{n}Y_{i}\right| \geq \varepsilon\right\} \leq 2 \exp\left\{-n(\varepsilon/2b)\left((1+\sigma^{2}/2b\varepsilon) \log (1+2b\varepsilon/\sigma^{2}) - 1\right)\right\}$$

$$\leq 2 \exp\left\{-n\varepsilon^{2}/2(\sigma^{2}+b\varepsilon)\right\},$$

where the latter inequality follows from the fact that log (1+u) > 2u/(2+u) for all u>0. Now, let $\epsilon>0$ be arbitrary and note that

$$|f_n(x) - f(x)| \le |f_n(x) - E\{f_n(x)\}| + |E\{f_n(x)\}| - f(x)|$$
.

From Lemma 2 (see Appendix) we know that $E\{Y_n\} \stackrel{\eta}{\to} f(x)$ if $Y_n = h_n^{-d} K((x-X_n)/h_n)$. Hence, by the Kronecker lemma [18, p. 238],

$$E\{f_n(x)\} = n^{-1} \sum_{i=1}^n E\{Y_i\} \stackrel{n}{\to} f(x)$$
.

Next, if $c = \sup_{i} E\{Y_i\}$ and $M = \sup_{x} K(x)$, then $E\{Y_i^2\} \le cMh_i^{-d}$ and $|Y_i - E\{Y_i\}| \le Mh_i^{-d}$. Therefore,

$$P\{|f_{n}(x) - E\{f_{n}(x)\}| \geq \epsilon\} = P\left\{\left|n^{-1}\sum_{i=1}^{n}(Y_{i}^{-E}\{Y_{i}^{-E}\})\right| \geq \epsilon\right\}$$

$$\leq 2 \exp\left\{-n\epsilon^{2}/2(cM\ell_{n}^{n} + M\epsilon\ell_{n}^{n})\right\},$$

where $\ell_n = \sup_{i \le n} h_i^{-d}$. The in probability part of Theorem 2 follows from Lemma 1 and (3). A well-known version of the strong law of large numbers (Loeve [18, p. 253]) asserts that if $|Y_n| \le Ln$ for all n and some $L < \infty$, (which is the case here since the h_n are positive numbers and $nh_n^d/\log\log n \xrightarrow{n} \infty$), then

$$n^{-1} \sum_{i=1}^{n} (Y_i - E\{Y_i\}) \stackrel{n}{\to} 0 \text{ w.p.1}$$

if and only if for every $\varepsilon > 0$

$$\sum_{k=0}^{\infty} P\left\{ \left| 2^{-k} \sum_{i=2^{k+1}}^{2^{k+1}} (Y_i - E\{Y_i\}) \right| \geq \epsilon \right\} < \infty$$

We thus conclude that $f_n(x) - E\{f_n(x)\} \stackrel{n}{\rightarrow} 0 \text{ w.p.1}$ if

$$\sum_{k=0}^{\infty} \left\{ 2 \exp -(2^k \varepsilon^2 / (2(cM+M\varepsilon)) \inf_{i \le 2^k} h_i^d \right\} < \infty$$

which is implied by

$$2^k \inf_{i \le 2^k} h_i^d / \log k \stackrel{k}{\to} \infty$$
.

This in turn follows from n inf $h_1^d/\log\log n \stackrel{n}{\to} \infty$ and $nh_n^d/\log\log n \stackrel{n}{\to} \infty \text{ (Lemma 1).}$

Remark 1. (Conditions for weak consistency). An application of Chebyshev's inequality allows a relaxation of condition (3) for weak consistency. We will see that it can be replaced by

(11)
$$n^{-2} \sum_{i=1}^{n} h_{i}^{-d} \stackrel{n}{\to} 0.$$

(Indeed, by Toeplitz's lemma [18, p. 238], $\mathrm{nh}_n^{d} \overset{n}{\to} \infty$ entails

$$n^{-2} \sum_{i=1}^{n} h_{i}^{-d} \le n^{-1} \sum_{i=1}^{n} (ih_{i})^{-d} \stackrel{n}{\Rightarrow} 0.$$

The example in which $h_n^d = n^{-\frac{1}{2}}$ except when $n = 2^k$, in which case $h_n^d = n^{-1}$, is one satisfying (11) but not (3). However, for monotone $\{h_n\}$ both conditions are equivalent in view of

$$(2n)^{-2} \sum_{i=1}^{2n} h_i^{-d} \ge (2n)^{-2} \sum_{i=1}^{2n} h_i^{-d} \ge (2n)^{-2} nh_n^{-d} = (4nh_n^d)^{-1}$$
.

They are also equivalent whenever (7) holds with $\,c$ = 1. To see this, assume that (11) is true but $\,nh_n^d \not\to \infty.\,$ In view of (7) this in turn implies that $\,nh_n^d \le M < \infty.\,$ But then

$$n^{-2} \sum_{i=1}^{n} h_{i}^{-d} \ge n^{-2} \sum_{i=1}^{n} i/M \ge n^{2}/(2n^{2}M) = 1/2M > 0$$
,

contradicting (11). We note that the weak consistency part of Theorem 2 under Condition A(i) can also be found in Ahmad and Lin [2].

Remark 2. (Conditions for strong consistency). The strong consistency of f_n was earlier established by Davies [8] under stricter conditions. Deheuvels [12] has shown that under weak additional assumptions on f, K, and $\{h_n\}$, (8) is necessary for the strong pointwise consistency of f_n . In the next theorem we will see that (8) can be replaced by

(12)
$$\sum_{n=1}^{\infty} n^{-2} h_n^{-d} < \infty$$

without compromising the strong pointwise consistency of f_n . The proof of this is based on Kolmogorov's second moment version of the strong law of large numbers. The sequence $\left\{ nh_n^d (\log n)/(\log \log n) \right\}$ is one satisfying (8) and (9) but not (12) since $\{1/n(\log n)(\log \log n)\}$ is not summable. On the contrary, the sequence

$$h_n^d = \begin{cases} n^{-\frac{1}{2}}, & n \neq 2^k \\ n^{-1}, & n = 2^k \end{cases}$$

satisfies (12) but not (3), and thus certainly not (6), (8), or (9). Thus the conditions (8) and (12) do not imply each other.

Suppose next that $\{h_n\}$ is sufficiently well-behaved so that Deheuvels' quasi-monotonicity condition (7) holds with c=1. With some work one can show that (7) and (12) imply (9):

$$nh_n^d/\log n \stackrel{n}{\to} \infty$$
.

Thus, for nicely behaved $\{h_n\}$ (e.g., nh_n^d is asymptotically monotone), Theorem 2 is stronger than the said theorem in which (12) replaces (8). (The proof of the latter implication uses the facts that for nondecreasing $\{\alpha_n\}$, $\sum 1/\alpha_n < \infty$ if and only if $\sum 2^n/\alpha_n < \infty$, and that $\sum 1/\alpha_n < \infty$ implies $n\alpha_n \stackrel{n}{\to} 0$, to conclude, upon taking $\alpha_n = nh_n^d$, that $\sum 1/(n\alpha_n) < \infty$ yields $\alpha_n/\log n \stackrel{n}{\to} \infty$.)

We now state the announced theorem.

THEOREM 3. If $\{h_n\}$ satisfies (2), (11), if K is a bounded density, and if Condition A holds, then $E\{(f_n(x)-f(x))^2\} \stackrel{n}{\to} 0$. If, additionally, (12) holds, then $f_n(x) \stackrel{n}{\to} f(x)$ w.p.1 as well.

Proof of Theorem 3. We need only show that $|f_n(x) - E\{f_n(x)\}| \stackrel{n}{\to} 0$ in the quadratic mean or with probability one. By Chebyshev's inequality and (11),

$$\begin{split} & \mathbb{P}\left\{\left|\,\mathbf{f}_{n}(\mathbf{x})\,-\,\mathbb{E}\left\{\mathbf{f}_{n}(\mathbf{x})\,\right\}\,\right| \,\,\geq\,\, \epsilon\,\right\} \leq\,\, \epsilon^{-2}\,\,\mathbb{E}\left\{\left(\mathbf{f}_{n}(\mathbf{x})\,-\,\mathbb{E}\left\{\mathbf{f}_{n}(\mathbf{x})\,\right\}\right)^{2}\right\} \\ & \leq\,\, \left(n\epsilon\right)^{-2}\,\,\sum_{i=1}^{n}\,\,\mathbb{E}\left\{\mathbf{Y}_{i}^{2}\right\} \,\,\leq\,\, \left(n\epsilon\right)^{-2}\,\,\sum_{i=1}^{n}\,\,\operatorname{cMh}_{i}^{-d}\,\stackrel{n}{\to}\,0\,\,\,. \end{split}$$

This proves the first part of Theorem 3. By Kolmogorov's condition for the strong law of large numbers (see [18, p. 253]) we know that $f_n(x) - E\{f_n(x)\} \stackrel{n}{\to} 0 \text{ w.p.1 if}$

$$\sum_{n=1}^{\infty} E\{Y_n^2\}/n^2 < \infty .$$

Using $\mathrm{E}\{Y_n^2\} \le \mathrm{cMh}_n^{-d}$, this condition reduces to (12). Actually, the inequality of Hajek and Renyi [17, pp. 258-260] provides us with an upper bound

$$P\left\{ \bigcup_{n=k}^{\infty} \left\{ \left| f_n(\mathbf{x}) - E\{f_n(\mathbf{x})\} \right| \geq \epsilon \right\} \right\} \leq (k\epsilon)^{-2} \sum_{n=1}^{k} E\{Y_n^2\} + \epsilon^{-2} \sum_{n=k+1}^{\infty} E\{Y_n^2\} / n^2$$

which tends to 0 as $k \to \infty$ for all $\epsilon > 0$.

Notice that f_n and g_n are valid densities on R^d . Thus, we can apply a theorem of Glick [16] which states that if $\{f_n\}$ is a sequence of densities converging to a density f in probability (w.p.1) for almost all x with respect to Lebesgue measure, and if the f_n are Borel measurable functions of x, X_1, \ldots, X_n , then $\int |f_n(x) - f(x)| dx \stackrel{n}{\to} 0$ in probability (w.p.1). Thus, the following theorem is a corollary to Theorems 2 and 3.

THEOREM 4. Let K be a bounded probability density and let any one of the following conditions hold:

- (i) f is almost everywhere continuous and (5) holds,
- (ii) f is bounded,
- (iii) K has compact support.

If (2) and (3) hold, then $\int |g_n(x) - f(x)| dx \stackrel{n}{\to} 0$ in probability. If also (6) holds, then $\int |g_n(x) - f(x)| dx \stackrel{n}{\to} 0$ w.p.1. If (2) and (11) hold, then $\int |f_n(x) - f(x)| dx \stackrel{n}{\to} 0$ in probability. If (2), and (8) or (12) hold, then $\int |f_n(x) - f(x)| dx \stackrel{n}{\to} 0$ w.p.1.

We emphasize that Theorem 4 holds for αll densities f if K has compact support, i.e., if for some finite M, K(x)=0 whenever ||x||>M. The part of Theorem 4 involving g_n was shown by Devroye and Wagner [15] for kernels K satisfying a condition weaker than (iii), namely

$$\int \sup_{y: ||y|| > ||x||} k(y) dx < \infty.$$

Remark 3. (Convergence in L_2 of density estimates). Because of its mathematical attractiveness, several authors consider $\int (f_n(x) - f(x))^2 dx$ as a global measure of the deviation of f_n from f. For instance, Wagner [26] shows that $\int (f_n(x) - f(x))^2 dx \stackrel{n}{=} 0$ w.p.l if K is a bounded density with $\int ||x|| |K(x)| dx < \infty$, if h_n is nonincreasing, if $\sum h_n/n < \infty$, if (12) holds, and if f satisfies a uniform Lipschitz condition.

A variety of results are available regarding $\int (g_n(x)-f(x))^2 dx$ (see Bickel and Rosenblatt [5], Ahmad [1] and Nadaraya [19]). Most of them are concerned with the rate of convergence to 0 of the mean integrated square error. It is known that if K *is a bounded density satisfying (5), if (2)-(3) hold, and if f^2 is integrable and f is almost everywhere continuous, then $E \left\{ \int (g_n(x) - f(x))^2 dx \right\} \stackrel{n}{\to} 0$ [13]. If in addition

$$nh_n^{2d}/log n \stackrel{n}{\rightarrow} \infty$$
,

Nadaraya [19] shows that $\int (g_n(x) - f(x))^2 dx \stackrel{\eta}{\to} 0$ w.p.l. (The latter condition is not necessary. Nadaraya's argument remains valid if we replace it by (9) and use Bennett's inequality [4] in the proper places.) In any case, all these results are only applicable if f^2 is integrable. Since f is unknown in advance this condition cannot be checked. Theorem 4 shows that the situation is not all that bad since the integrated L_1 error converges to 0 w.p.l for all densities f (under mild conditions on the sequence $\{h_n\}$ which we can pick anyway) and all the estimates discussed in this paper. This result is not surprising since L_1 , not L_2 , is the natural space in which the properties of densities should be studied. It is comforting to know that the requirement that f be almost everywhere continuous can be dropped altogether.

3. A Note on the Deheuvels Estimator.

In this section we complement some results of Deheuvels [10-12].

THEOREM 5. If $\{h_n^{}\}$ satisfies (10), if K is a bounded probability density and if Condition A holds, then $\bar{f}_n(x) \stackrel{n}{\to} f(x)$ w.p.1.

Proof of Theorem 5. Notice first that

$$\mathbb{E}\left\{\overline{f}_{n}(x)\right\} = \left(\sum_{i=1}^{n} h_{i}^{d}\right)^{-1} \sum_{i=1}^{n} h_{i}^{d} \mathbb{E}\left\{Y_{i}\right\} ,$$

where Y is defined as in the proof of Theorem 2. Using Lemma 2 $(\text{E}\{Y_n\}^{\frac{n}{4}}f(x)) \text{ and Toeplitz's lemma [18, p. 238] we see that } \Sigma \ h_n^d = \infty$ implies that $\text{E}\{\overline{f}_n(x)\}^{\frac{n}{4}}f(x)$. Thus we need only show that $\overline{f}_n(x) - \text{E}\{\overline{f}_n(x)\}^{\frac{n}{4}} \text{ 0 w.p.1.} \text{ By Loeve's criterion for the strong law of large numbers [18, p. 253] this follows if}$

$$\sum_{n=1}^{\infty} E\{K^{2}((x-X_{n})/h_{n})\} / \left(\sum_{i=1}^{n} h_{i}^{d}\right)^{2} < \infty .$$

But we know that $E\{K^2((x-X_n)/h_n)\} \le Mh_n^d E\{Y_n\} \le Mch_n^d$. Theorem 5 now follows since $\sum_{n=1}^{\infty} h_n^d = \infty$ implies

$$\sum_{n=1}^{\infty} \ \mathbf{h}_{n}^{d} \!\! \left(\!\! \left(\sum_{i=1}^{n} \ \mathbf{h}_{i}^{d} \right)^{\!\! 2} \right. < \infty$$
 .

To see this, assume that $h_1 > 0$ and notice that

$$\begin{split} & \sum_{n=1}^{\infty} \ h_{n}^{d} / \left(\sum_{i=1}^{n} \ h_{i}^{d} \right)^{2} \leq 1/h_{1}^{d} + \sum_{n=2}^{\infty} \ h_{n}^{d} / \left(\left(\sum_{i=1}^{n} \ h_{i}^{d} \right) \left(\sum_{i=1}^{n-1} \ h_{i}^{d} \right) \right) \\ & \leq \sum_{n=2}^{\infty} \left(1 / \left(\sum_{i=1}^{n-1} \ h_{i}^{d} \right) - 1 / \left(\sum_{i=1}^{n} \ h_{i}^{d} \right) \right) + 1/h_{1}^{d} \\ & = 2/h_{1}^{d} \ . \end{split}$$

THEOREM 6. If $\{h_n^{}\}$ satisfies (10), if K is a bounded probability density and if any one of the following conditions hold,

- (i) f is almost everywhere continuous and (5) is satisfied,
- (ii) f is bounded,
- (iii) K has compact support,

then $\int |\overline{f}_n(x) - f(x)| dx \stackrel{n}{\to} 0 w.p.1.$

One should notice that (10) is weaker than (11) since

$$\sum_{i=1}^{n} h_{i}^{d} \ge n^{2} \left(\sum_{i=1}^{n} h_{i}^{-d} \right)^{-1} \stackrel{n}{\Rightarrow} 0.$$

Here we used the following inequality valid for all sequences $\{h_n^{}\}:$

$$\left(n^{-1}\sum_{i=1}^{n} h_{i}^{-d}\right)\left(n^{-1}\sum_{i=1}^{n} h_{i}^{d}\right) \geq 1$$
.

The conditions (10) cannot be improved upon in Theorem 5 due to a result by Deheuvels [12] who shows that under some additional restrictions on f, K, and $\{h_n\}$, in order for $\overline{f}_n(x) \stackrel{n}{\to} f(x)$ w.p.l, it is necessary that (10) hold.

Proof of Theorem 6. Theorem 6 follows from Theorem 5, Glick's result [16], and the fact that almost all x are Lebesgue points for f.

LEMMA 1. If $a_n, b_n \ge 0$, $a_n + \infty$, then $a_n/b_n \xrightarrow{n} \infty$ if and only if $a_n/(\sup_{i \le n} b_i) \xrightarrow{n} \infty$.

Proof. Notice that

$$a_n/b_n \ge a_n/(\sup_{i\le n} b_i) \ge \min(\inf_{i>N} a_i/b_i, a_n/(\sup_{i\le N} b_i))$$
.

Lemma 1 follows by first picking $\,N\,$ large enough and then letting $\,n\,$ grow unbounded.

COROLLARY. The conditions $\operatorname{nh}_n^{d} \stackrel{n}{\to} \infty$ and $\operatorname{n} \inf_{i \leq n} \operatorname{h}_i^{d} \stackrel{n}{\to} \infty$ are equivalent. Furthermore, the conditions $\operatorname{nh}_n^d/\log\log n \stackrel{n}{\to} \infty$ and $\operatorname{n} \inf_{i \leq n} \operatorname{h}_i^d/\log\log n \stackrel{n}{\to} \infty$ are equivalent.

LEMMA 2. Let K be a bounded probability density, let Condition A hold, and let $h_n \stackrel{n}{\to} 0$. Then

$$\mathbb{E}\left\{h_n^{-d} \ \mathbb{K}((x-X_1)/h_n)\right\} \overset{n}{\to} f(x) \ .$$

Proof. We need only show that

(13)
$$\left| \int h_n^{-d} K((x-y)/h_n) f(y) dy - f(x) \right| \\ \leq \int h_n^{-d} K((x-y)/h_n) \left| f(y) - f(x) \right| dy \stackrel{n}{\to} 0.$$

Let $S(\alpha,\rho)$ be a closed sphere in R^d centered at α with radius ρ and let $(\cdot)^C$ denote the complement of a set. We can upper bound (13) by

$$\begin{split} \left(\rho h_n\right)^{-d} & \int\limits_{S(x,\rho h_n)} M \big| f(y) - f(x) \big| dy \\ & + \min(\sup\limits_{S^c(x,\rho)} K(y); (\sup\limits_{S^c(x,\rho)} f(y)) & \int\limits_{S^c(x,\rho)} K(y) dy) \ , \end{split}$$

where M = sup K(y) and ρ is chosen. The last term can be made small by choice of ρ if either K has compact support or f is bounded. The first term is small for large n if $h_n \stackrel{n}{\to} 0$ and x is a Lebesgue point of f. Expression (13) can also be upper bounded by

$$\sup_{S(\mathbf{x},\rho)} \left| f(y) - f(\mathbf{x}) \right| + f(\mathbf{x}) \int_{S^{c}(\mathbf{x},\rho/h_{n})} K(y) \, \mathrm{d}y + \rho^{-d} \sup_{S^{c}(\mathbf{x},\rho/h_{n})} \left| \left| y \right| \right|^{d} K(y) \ .$$

If x is a continuity point of f, then we can make the first term small by choice of ρ . The last two terms are small for large n if (5) holds, if K is integrable, and if $h_n \stackrel{\eta}{\to} 0$.

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